

**Applications of the large deviation technique**

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**1 Safety capital for an insurance company**

Consider some insurance company which needs to decide on the amount of capital  $S_0$  it needs to hold to avoid the cash flow issue. Suppose the insurance premium per month is a fixed (non-random) quantity  $C > 0$ . Suppose the claims are i.i.d. random variable  $A_N \geq 0$  for the time  $N = 1, 2, \dots$ . Then the capital at time  $N$  is  $S_N = S_0 + \sum_{n=1}^N (C - A_n)$ . The company wishes to avoid the situation where the cash flow  $S_N$  is negative. Thus it needs to decide on the capital  $S_0$  so that  $\mathbb{P}(\exists N, S_N \leq 0)$  is small. Obviously this involves a tradeoff between the "smallness" and the amount  $S_0$ . Let us assume that upper bound  $\delta = 0.001$ , namely 0.1% is acceptable (in fact this is pretty close to the banking regulation standards). We have

$$\begin{aligned} \mathbb{P}(\exists N, S_N \leq 0) &= \mathbb{P}(\min_N S_0 + \sum_{n=1}^N (C - A_n) \leq 0) \\ &= \mathbb{P}(\max_N \sum_{n=1}^N (A_n - C) \geq S_0) \end{aligned}$$

If  $\mathbb{E}[A_1] \geq C$ , we have  $\mathbb{P}(\max_N \sum_{n=1}^N (A_n - C) \geq S_0) = 1$ . Thus, the interesting case is  $\mathbb{E}[A_1] < C$  (negative drift), and the goal is to determine the starting capital  $S_0$  such that

$$\mathbb{P}(\max_N \sum_{n=1}^N (A_n - C) \geq S_0) \leq \delta.$$

## 2 Buffer overflow in a queueing system

The following model is a variant of a classical so called GI/GI/1 queueing system. In application to communication systems this queueing system consists of a single server, which processes some  $C > 0$  number of communication packets per unit of time. Here  $C$  is a fixed deterministic constant. Let  $A_n$  be the random number packets arriving at time  $n$ , and  $Q_n$  be the queue length at time  $n$  (assume  $Q_0=0$ ). By recursion, we have that

$$\begin{aligned} Q_N &= \max(Q_{N-1} + A_N - C, 0) \\ &= \max(Q_{N-2} + A_{N-1} + A_N - 2C, A_N - C, 0) \\ &= \max_{1 \leq n \leq N-1} \left( \sum_{k=1}^n (A_{N-k} - C), 0 \right) \end{aligned}$$

Notice, that in distributional sense we have

$$Q_N = \max \left( \max_{1 \leq n \leq N-1} \left( \sum_{k=1}^n (A_{N-k} - C) \right), 0 \right)$$

In steady state, i.e.  $N = \infty$ , we have

$$Q_\infty = \max \left( \max_{n \geq 1} \sum_{k=1}^n (A_k - C), 0 \right)$$

Our goal is to design the size of the queue length storage (buffer)  $B$ , so that the likelihood that the number of packets in the queue exceeds  $B$  is small. In communication application this is important since every packet not fitting into the buffer is dropped. Thus the goal is to find buffer size  $B > 0$  such that

$$\mathbb{P}(Q_\infty \geq B) \leq \delta \Rightarrow \mathbb{P}(\max_{n \geq 1} \sum_{k=1}^n (A_k - C) \geq B) \leq \delta$$

If  $\mathbb{E}[A_1] \geq C$ , we have  $\mathbb{P}(Q_\infty \geq B) = 1$ . So the interesting case is  $\mathbb{E}[A_1] < C$  (negative drift).

## 3 Buffer overflow probability

We see that in both situations we need to estimate

$$\mathbb{P}(\max_{n \geq 1} \sum_{k=1}^n (A_k - C) \geq B).$$

We will do this asymptotically as  $B \rightarrow \infty$ .

**Theorem 1.** Given an i.i.d. sequence  $A_n \geq 0$  for  $n \geq 1$  and  $C > \mathbb{E}[A_1]$ .  
 Suppose

$$M(\theta) = \mathbb{E}[\exp(\theta A)] < \infty, \text{ for some } \theta \in [0, \theta_0).$$

Then

$$\lim_{B \rightarrow \infty} \frac{1}{B} \log \mathbb{P}(\max_{n \geq 1} \sum_{k=1}^n (A_k - C) \geq B) = -\sup\{\theta > 0 : M(\theta) < \exp(\theta C)\}$$

Observe that since  $A_n$  is non-negative, the MGF  $\mathbb{E}[\exp(\theta A_n)]$  is finite for  $\theta < 0$ . Thus it is finite in an interval containing  $\theta = 0$ , and applying the result of Lecture 2 we can take the derivative of MGF. Then

$$\left. \frac{d}{d\theta} M(\theta) \right|_{\theta=0} = \mathbb{E}[A], \quad \left. \frac{d}{d\theta} \exp(\theta C) \right|_{\theta=0} = C$$

Since  $\mathbb{E}[A_n] < C$ , then there exists small enough  $\theta$  so that  $M(\theta) < \exp(\theta C)$ ,

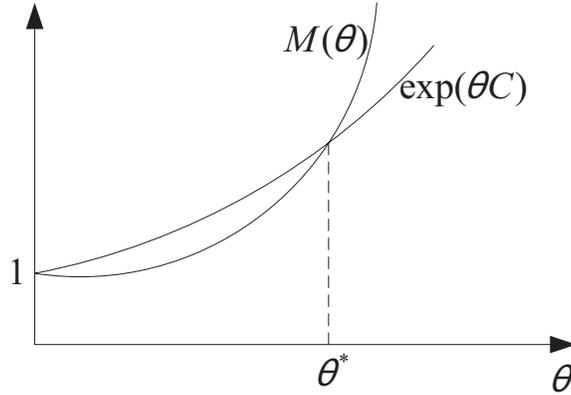


Figure 1: Illustration for the existence of  $\theta$  such that  $M(\theta) < \exp(\theta C)$

and thus the set of  $\theta > 0$  for which this is the case is non-empty. (see Figure 1).  
 The theorem says that roughly speaking

$$\mathbb{P}(\max_n \sum_{k=1}^n (A_k - C) \geq B) \sim \exp(-\theta^* B),$$

when  $B$  is large. Thus given  $\delta$  select  $B$  such that  $\exp(-\theta^* B) \leq \delta$ , and we can set  $B = \frac{1}{\theta^*} \log \frac{1}{\delta}$ .

Example. Let  $A$  be a random variable uniformly distributed in  $[0, a]$  and  $C = 2$ . Then, the moment generating function of  $A$  is

$$M(\theta) = \int_0^a \exp(\theta t) a^{-1} dt = \frac{\exp(\theta a) - 1}{\theta a}$$

Then

$$\sup\{\theta > 0 : M(\theta) \leq \exp(\theta C)\} = \sup\{\theta > 0 : \frac{\exp(\theta a) - 1}{\theta a} \leq \exp(2\theta)\}$$

Case 1:  $a = 3$ , we have  $\theta^* = \sup\{\theta > 0 : \exp(3\theta) - 1 \leq 3\theta \exp(2\theta)\}$ , i.e.  $\theta^* = 1.54078$ .

Case 2:  $a = 4$ , we have that  $\{\theta > 0 : \exp(3\theta) - 1 \leq 3\theta \exp(2\theta)\} = \emptyset$  since  $\mathbb{E}[A] = 2 = C$ .

Case 3:  $a = 2$ , we have that  $\{\theta > 0 : \exp(3\theta) - 1 \leq 3\theta \exp(2\theta)\} = \mathbb{R}_+$  and thus  $\theta^* = \infty$ , which implies that  $\mathbb{P}(\max_n \sum_{k=1}^n (A_k - C) \geq B) = 0$  by theorem 1.

*Proof of Theorem 1.* We will first prove an upper bound and then a lower bound. Combining them yields the result. For the upper bound, we have that

$$\begin{aligned} \mathbb{P}(\max_n \sum_{k=1}^n (A_k - C) \geq B) &\leq \sum_{n=1}^{\infty} \mathbb{P}(\sum_{k=1}^n (A_k - C) \geq B) \\ &= \sum_{n=1}^{\infty} \mathbb{P}(\frac{1}{n} \sum_{k=1}^n A_k \geq C + \frac{B}{n}) \\ &\leq \sum_{n=1}^{\infty} \exp(-n(\theta(C + \frac{B}{n}) - \log M(\theta))) \quad (\theta > 0) \\ &= \exp(-\theta B) \sum_{n \geq 1} \exp(-n(\theta C - \log M(\theta))) \end{aligned}$$

Fix any  $\theta$  such that  $\theta C \geq \log M(\theta)$ , the inequality above gives

$$\begin{aligned} &\leq \exp(-\theta B) \sum_{n \geq 0} \exp(-n(\theta C - \log M(\theta))) \\ &= \exp(-\theta B) [1 - \exp(-(\theta C - \log M(\theta)))]^{-1} \end{aligned}$$

Simplification of the inequality above gives

$$\begin{aligned} \frac{1}{B} \log \mathbb{P}(\max_n \sum_{k=1}^n (A_k - C) \geq B) &\leq -\theta + \frac{1}{B} \log([1 - \exp(-(\theta C - \log M(\theta)))]^{-1}) \\ \Rightarrow \limsup_{B \rightarrow \infty} \frac{1}{B} \log \mathbb{P}(\max_n \sum_{k=1}^n (A_k - C) \geq B) &\leq -\theta \text{ for } \forall \theta : M(\theta) < \exp(\theta C) \end{aligned}$$

Next, we will derive the lower bound.

$$\mathbb{P}(\max_n \sum_{k=1}^n (A_k - C) \geq B) \geq \mathbb{P}(\sum_{k=1}^n (A_k - C) \geq B), \forall n$$

Fix a  $t > 0$ , then

$$\begin{aligned} \mathbb{P}(\max_n \sum_{k=1}^n (A_k - C) \geq B) &\geq \mathbb{P}(\sum_{k=1}^{\lceil Bt \rceil} (A_k - C) \geq B) \\ &\geq \mathbb{P}\left(\sum_{k=1}^{\lceil Bt \rceil} (A_k - C) \geq \frac{\lceil Bt \rceil}{t}\right) \end{aligned}$$

Then, we have

$$\begin{aligned} &\liminf_B \frac{1}{B} \log \mathbb{P}(\max_n \sum_{k=1}^n (A_k - C) \geq B) \\ &\geq \liminf_B \frac{1}{B} \log \mathbb{P}\left(\sum_{k=1}^{\lceil Bt \rceil} (A_k - C) \geq \frac{\lceil Bt \rceil}{t}\right) \\ &= \liminf_B \frac{t}{\lceil Bt \rceil} \log \mathbb{P}\left(\sum_{k=1}^{\lceil Bt \rceil} (A_k - C) \geq \frac{\lceil Bt \rceil}{t}\right) \\ &= t \liminf_n \frac{1}{n} \log \mathbb{P}\left(\sum_{k=1}^n (A_k - C) \geq \frac{n}{t}\right) \\ &= t \liminf_n \frac{1}{n} \log \mathbb{P}\left(\frac{1}{n} \sum_{k=1}^n A_k \geq C + \frac{1}{t}\right) \\ &\geq -t \inf_{x > C + \frac{1}{t}} I(x) \quad (\text{by Cramer's theorem.}) \\ &\geq -\inf_{t > 0} t \inf_{x > C + \frac{1}{t}} I(x) \quad (\text{since we can choose an arbitrary positive } t.) \quad (1) \end{aligned}$$

We claim that

$$-\inf_{t > 0} t \inf_{x > C + \frac{1}{t}} I(x) = -\inf_{t > 0} t I(C + \frac{1}{t})$$

Indeed, let  $x^* = \inf x : I(x) = \infty$  (possibly  $x^* = \infty$ ). If  $x^* \leq C$ , then  $I(C + \frac{1}{t}) = \infty$ . Suppose  $C < x^*$ . If  $t$  is such that  $C + \frac{1}{t} \geq x^*$ , then  $\inf_{x > C + \frac{1}{t}} I(x) = \infty$ . Therefore it does not make sense to consider such  $t$ . Now for  $c + \frac{1}{t} < x^*$ ,

we have  $I$  is convex non-decreasing and finite on  $[E[A_1], x^*)$ . Therefore it is continuous on  $[E[A_1], x^*)$ , which gives that

$$\inf_{x > C + \frac{1}{t}} I(x) = I\left(C + \frac{1}{t}\right)$$

and the claim follows. Thus, we obtain

$$\liminf_{B \rightarrow \infty} \frac{1}{B} \log \mathbb{P}\left(\max_n \sum_{k=1}^n (A_k - C) \geq B\right) \geq -\inf_{t > 0} tI\left(C + \frac{1}{t}\right)$$

Exercise in HW 2 shows that  $\sup\{\theta > 0 : M(\theta) < \exp(C\theta)\} = \inf_{t > 0} tI\left(C + \frac{1}{t}\right)$ .

□

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