

OLS: Estimation and Standard Errors

Brandon Lee

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Ordinary Least Squares

- The model:

$$y = X\beta + \varepsilon$$

where y and ε are column vectors of length n (the number of observations), X is a matrix of dimensions n by k (k is the number of parameters), and β is a column vector of length k .

- For every observation $i = 1, 2, \dots, n$, we have the equation

$$y_i = x_{i1}\beta_1 + \dots + x_{ik}\beta_k + \varepsilon_i$$

- Roughly speaking, we need the orthogonality condition

$$E[\varepsilon_i x_i] = 0$$

for the OLS to be valid (in the sense of consistency).

- We want to find $\hat{\beta}$ that solves

$$\min_{\beta} (y - X\beta)'(y - X\beta)$$

- The first order condition (in vector notation) is

$$0 = X'(y - X\hat{\beta})$$

and solving this leads to the well-known OLS estimator

$$\hat{\beta} = (X'X)^{-1} X'y$$

Geometric Interpretation

- The left-hand variable is a vector in the n -dimensional space. Each column of X (regressor) is a vector in the n -dimensional space as well, and we have k of them. Then the subspace spanned by the regressors forms a k -dimensional subspace of the n -dimensional space. The OLS procedure is nothing more than finding the orthogonal projection of y on the subspace spanned by the regressors, because then the vector of residuals is orthogonal to the subspace and has the minimum length.
- This interpretation is very important and intuitive. Moreover, this is a unique characterization of the OLS estimate.
- Let's see how we can make use of this fact to recognize OLS estimators in disguise as more general GMM estimators.

Interest Rate Model

- Refer to pages 35-37 of Lecture 7.
- The model is

$$r_{t+1} = a_0 + a_1 r_t + \varepsilon_{t+1}$$

where

$$E[\varepsilon_{t+1}] = 0$$

$$E[\varepsilon_{t+1}^2] = b_0 + b_1 r_t$$

- One easy set of moment conditions:

$$0 = E[(1, r_t)' (r_{t+1} - a_0 - a_1 r_t)]$$

$$0 = E[(1, r_t)' \left((r_{t+1} - a_0 - a_1 r_t)^2 - b_0 - b_1 r_t \right)]$$

- Solving these sample moment conditions for the unknown parameters is exactly equivalent to a two-stage OLS procedure.
- Note that the first two moment conditions give us

$$E_T [(1, r_t)' (r_{t+1} - \hat{a}_0 - \hat{a}_1 r_t)] = 0$$

But this says that the estimated residuals are orthogonal to the regressors and hence \hat{a}_0 and \hat{a}_1 must be OLS estimates of the equation

$$r_{t+1} = a_0 + a_1 r_t + \varepsilon_{t+1}$$

- Now define

$$\hat{\varepsilon}_{t+1} = r_{t+1} - \hat{a}_0 - \hat{a}_1 r_t$$

then the sample moment conditions

$$E_T \left[(1, r_t)' \left((r_{t+1} - \hat{a}_0 - \hat{a}_1 r_t)^2 - \hat{b}_0 - \hat{b}_1 r_t \right) \right] = 0$$

tell us that \hat{b}_0 and \hat{b}_1 are OLS estimates from the equation

$$\hat{\varepsilon}_{t+1}^2 = b_0 + b_1 r_t + u_{t+1}$$

by the same logic.

Standard Errors

- Let's suppose that $E[\varepsilon_i^2|X] = \sigma^2$ and $E[\varepsilon_i\varepsilon_j|X] = 0$ for $i \neq j$. In other words, we are assuming independent and homoskedastic errors.
- What is the standard error of the OLS estimator under this assumption?

$$\begin{aligned} \text{Var}(\hat{\beta}|X) &= \text{Var}(\hat{\beta} - \beta|X) \\ &= \text{Var}\left((X'X)^{-1}X'\varepsilon|X\right) \\ &= (X'X)^{-1}X'\text{Var}(\varepsilon|X)X(X'X)^{-1} \end{aligned}$$

- Under the above assumption,

$$\text{Var}(\varepsilon|X) = \sigma^2 I_n$$

and so

$$\text{Var}(\hat{\beta}|X) = \sigma^2 (X'X)^{-1}$$

- We can estimate σ^2 by

$$\widehat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n \hat{\varepsilon}_i^2$$

and the standard error for the OLS estimator is given by

$$\widehat{\text{Var}}(\hat{\beta}|X) = \widehat{\sigma}^2 (X'X)^{-1}$$

- This is the standard error that most (less sophisticated) statistical softwares report.
- But it is rarely the case that it is safe to assume independent homoskedastic errors. The Newey-West procedure is a straightforward and robust method of calculating standard errors in more general situations.

Newey-West Standard Errors

- Again,

$$\begin{aligned}\text{Var}(\hat{\beta}|X) &= \text{Var}(\hat{\beta} - \beta|X) \\ &= \text{Var}\left((X'X)^{-1}X'\varepsilon|X\right) \\ &= (X'X)^{-1}\text{Var}(X'\varepsilon|X)(X'X)^{-1}\end{aligned}$$

- The Newey-West procedure boils down to an alternative way of looking at $\text{Var}(X'\varepsilon|X)$.
- If we suspect that the error terms may be heteroskedastic, but still independent, then

$$\widehat{\text{Var}}(X'\varepsilon|X) = \sum_{i=1}^n \hat{\varepsilon}_i^2 \cdot x_i x_i'$$

and our standard error for the OLS estimate is

$$\widehat{\text{Var}}(\hat{\beta}|X) = (X'X)^{-1} \left(\sum_{i=1}^n \hat{\varepsilon}_i^2 \cdot x_i x_i' \right) (X'X)^{-1}$$

- If we suspect correlation between error terms as well as heteroskedasticity, then

$$\widehat{Var}(X'\varepsilon|X) = \sum_{j=-k}^k \frac{k-|j|}{k} \left(\sum_{t=1}^n \hat{\varepsilon}_t \hat{\varepsilon}_{t+j} \cdot x_t x'_{t+j} \right)$$

and our standard error for the OLS estimator is

$$\widehat{Var}(\hat{\beta}|X) = (X'X)^{-1} \left(\sum_{j=-k}^k \frac{k-|j|}{k} \left(\sum_{t=1}^n \hat{\varepsilon}_t \hat{\varepsilon}_{t+j} \cdot x_t x'_{t+j} \right) \right) (X'X)^{-1}$$

- We can also write these standard errors to resemble the general GMM standard errors (see page 23 of Lecture 8).
- In the uncorrelated errors case, we have

$$\begin{aligned}
 \widehat{\text{Var}}(\hat{\beta}|X) &= (X'X)^{-1} \left(\sum_{i=1}^n \hat{\varepsilon}_i^2 \cdot x_i x_i' \right) (X'X)^{-1} \\
 &= \frac{1}{n} \left(\frac{X'X}{n} \right)^{-1} \left(\frac{1}{n} \sum_{i=1}^n \hat{\varepsilon}_i^2 \cdot x_i x_i' \right) \left(\frac{X'X}{n} \right)^{-1} \\
 &= \frac{1}{n} \hat{E}(x_i x_i')^{-1} \left(\frac{1}{n} \sum_{i=1}^n \hat{\varepsilon}_i^2 \cdot x_i x_i' \right) \hat{E}(x_i x_i')^{-1}
 \end{aligned}$$

and for the general Newey-West standard errors, we have

$$\begin{aligned}
 \widehat{\text{Var}}(\hat{\beta}|X) &= (X'X)^{-1} \left(\sum_{j=-k}^k \frac{k-|j|}{k} \left(\sum_{t=1}^n \hat{\varepsilon}_t \hat{\varepsilon}_{t+j} \cdot x_t x_{t+j}' \right) \right) (X'X)^{-1} \\
 &= \frac{1}{n} \hat{E}(x_i x_i')^{-1} \left(\frac{1}{n} \sum_{j=-k}^k \frac{k-|j|}{k} \left(\sum_{t=1}^n \hat{\varepsilon}_t \hat{\varepsilon}_{t+j} \cdot x_t x_{t+j}' \right) \right) \hat{E}(x_i x_i')^{-1}
 \end{aligned}$$

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